

# The Sufficiency Principle

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# Outline

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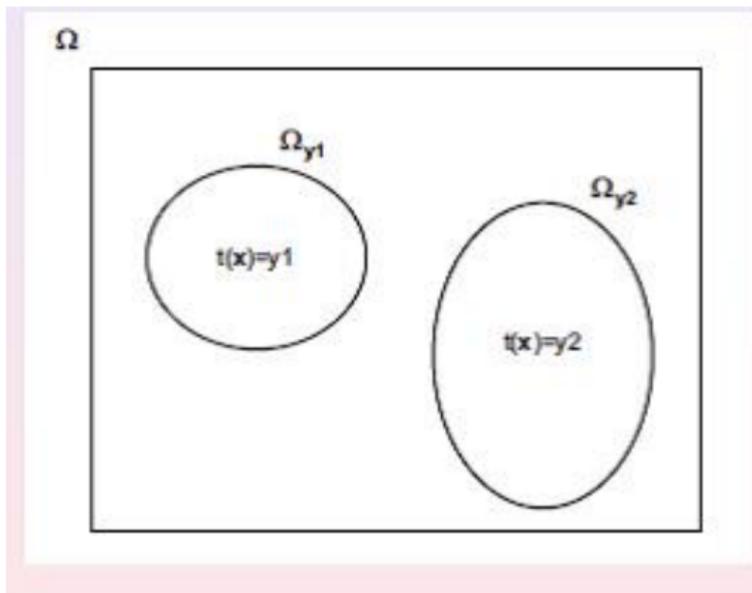
# Principle of Data Reduction

*"... we are suffering from a plethora of surmise, conjecture and hypothesis. The difficulty is to detach the framework of fact - of absolutely undeniable fact - from the embellishments of theorists and reporters."*

**Sherlock Holmes**  
Silver Blaze

# Sufficiency

Any sample statistic  $Y = t(\mathbf{X})$  provides a partition of the sample space into the subsets  $\Omega_y$ , where  $\Omega_y = \{\mathbf{x} : t(\mathbf{x}) = y\}$



# Sufficiency

- If two observed samples, say  $x_1$  and  $x_2$ , belong  $\Omega_Y$ , we have that  $t(x_1) = t(x_2)$  and they lead to the same inferential conclusion on  $\theta$ , even if the two samples are different.
- Any sample statistic defines a form of data reduction. The problem is that using a sample statistic we can lose information on  $\theta$ . It is then important to use a sample statistic that captures all the information about  $\theta$  contained in the sample, in the sense that it is equivalent to the sample about the information on  $\theta$ .

# Principles of data reduction

## Remarks

- If the sample size  $n$  is large the information in  $\mathbf{x} = (x_1, \dots, x_n)$  on  $\theta$  may be hard to understand
- The information in a sample can be summarized using statistics (such as sample mean...)
- Any sample statistic  $T(\mathbf{x})$  defines a form of data reduction or data summary
- The value  $t = T(\mathbf{x})$  instead of the entire sample  $\mathbf{x}$  is used to make inference on  $\theta$

The problem is that using a sample statistic we can lose information on  $\theta$ . It is then important to use a sample statistic that captures all the information about  $\theta$  contained in the sample, in the sense that it is equivalent to the sample about the information on  $\theta$ .

# Sufficiency

**SUFFICIENCY PRINCIPLE:** If  $T(X)$  is a sufficient statistic for  $\theta$ , then any inference about  $\theta$  should depend on the sample  $X$  only through the value  $T(X)$ . That is, if  $x$  and  $y$  are two points such that  $T(x) = T(y)$ , then any inference about  $\theta$  should be the same whether  $X = x$  or  $X = y$  is observed.

**Definition** A statistic  $T(X)$  is a *sufficient statistic* for  $\theta$  if the conditional distribution of the sample  $X$  given the value of  $T(X)$  does not depend on  $\theta$ .

# Sufficiency

Let  $T(X)$  be a sufficient statistic. Consider the pair  $(X, T(X))$ . Obviously,  $(X, T(X))$  contains the same information about  $\theta$  as  $X$  alone, since  $T(X)$  is a function of  $X$ . But if we know  $T(X)$ , then  $X$  itself has no values for us since its conditional distribution given  $T(X)$  is independent of  $\theta$ . Thus, by observing  $X$  (in addition to  $T(X)$ ), we cannot say whether one particular value of parameter  $\theta$  is more likely than another. Therefore, once we know  $T(X)$  we can discard  $X$  completely.

# Sufficiency

In other words, given the value  $t$  of a sufficient statistic  $T$ , conditionally there is no more **information** left in the original data regarding the unknown parameter  $\theta$ . Put another way, we may think of  $X$  trying to tell us a story about  $\theta$ , but once a sufficient summary  $T$  becomes available, the original story then becomes redundant. Observe that the whole data  $X$  is always sufficient for  $\theta$  in this sense. But, we are aiming at a *shorter* summary statistic which has the same amount of information available in  $X$ . Thus, once we find a sufficient statistic  $T$ , we will focus only on the summary statistic  $T$ .

## Example

Let  $X = (X_1, \dots, X_n)$  be a random sample from  $N(\mu, \sigma^2)$ . Suppose that  $\sigma^2$  is known. Let consider  $T(X) = \bar{X}_n$  and we know  $\bar{X}_n \sim N(\mu, \sigma^2/n)$ . Let's now calculate the conditional distribution of  $X$  given  $T(X) = t$ .

$$\begin{aligned} f_{X|T(X)}(x|T(X) = T(x)) &= \frac{f_X(x)}{f_T(T(x))} \\ &= \frac{\exp\left(-\sum_{i=1}^n (x_i - \mu)^2 / (2\sigma^2)\right)}{(2\pi)^{n/2} \sigma^n} \\ &= \frac{\exp\left(-n(\bar{x}_n - \mu)^2 / (2\sigma^2)\right)}{(2\pi)^{1/2} (\sigma/n)^{1/2}} \\ &= \frac{\exp\left(-\sum_{i=1}^n (x_i - \bar{x}_n)^2 / (2\sigma^2)\right)}{(2\pi)^{(n-1)/2} \sigma^{n-1} / n^{1/2}} \end{aligned}$$

and this is independent on  $\mu$ !  $\bar{X}$  is a sufficient statistic for  $\mu$ .

## Example

Let  $X$  be a random sample from  $Exp(\lambda)$ , let's consider the statistic  $T = I(X > 2)$ .

$$\begin{aligned}P(X > 3|T) &= \frac{P(X > 3 \cap T = 1)}{P(T = 1)} \\&= \frac{P(X > 3 \cap X > 2)}{P(X > 2)} \\&= \frac{P(X > 3)}{P(X > 2)} \\&= \frac{\exp(-3\lambda)}{\exp(-2\lambda)} \\&= \exp(-\lambda)\end{aligned}$$

The statistics  $T = I(X > 2)$  is NOT a sufficient statistics for  $\lambda$ .

# Sufficiency

**Theorem** If  $p(x|\theta)$  is the joint probability distribution function of  $X$ , and  $q(t|\theta)$  is the probability distribution function of  $T(X)$ , then  $T(X)$  is a sufficient statistic for  $\theta$  if, and only if, for every  $x$  in the sample space the ratio  $p(x|\theta)/q(T(x)|\theta)$  is constant as a function of  $\theta$ .

**Note:**  $p(x|\theta)$  indicates either discrete or continuous distribution (in this case  $p(x|\theta) = f(x|\theta)$ )

# Sufficiency

**Factorization Theorem** Let  $p(x|\theta)$  denote the joint probability distribution function of a sample  $X$ . A statistic  $T(X)$  is a sufficient statistics for  $\theta$  if and only if there exists a function  $g(t|\theta)$  and  $h(x)$  such that, for all sample points  $x$  and all parameter points  $\theta$ :

$$p(x|\theta) = g(T(x)|\theta)h(x).$$

*Proof* See Casella and Berger pag. 276

# Factorization Theorem

*Brief proof for more details see See Casella and Berger pag. 276*

Let's consider only the discrete case

$$\begin{aligned}P_{\theta}(X = x) &= \sum_t P_{\theta}(X = x \cap T = t) \\&= P_{\theta}(X = x \cap T = T(x)) \\&= P_{\theta}(X = x | T = T(x))P_{\theta}(T = T(x))\end{aligned}$$

Since  $T$  is sufficient  $P_{\theta}(X = x | T = T(x))$  is independent of  $\theta$  and so it is equal to  $h(x)$  and  $P_{\theta}(X = x) = g(T(x)|\theta)h(x)$

# Factorization Theorem

*Brief proof for more details see See Casella and Berger pag. 276*

On the other hand...

$$\begin{aligned} P_{\theta}(X = x | T = T(x)) &= \frac{P_{\theta}(X = x)}{P_{\theta}(T = T(x))} \\ &= \frac{g(T(x)|\theta)h(x)}{\sum_{T(y)=t} g(T(y)|\theta)h(y)} \end{aligned}$$

Since if  $T(x) \neq t$  then  $P_{\theta}(X = x | T = T(x)) = 0$ ,  $g(T(y)|\theta)$  is a constant

$$P_{\theta}(X = x | T = T(x)) = \frac{h(x)}{\sum_{T(y)=t} h(y)}$$

does not depend on  $\theta$

# Sufficiency and exponential family

- As a consequence of the factorization Theorem, we have that when a sample statistic, say  $s(\mathbf{X})$ , is a one-to-one function of a sufficient statistic for  $\theta$ , also  $s(\mathbf{X})$  is a sufficient statistic.
- When we sample from an exponential family, we have that

$$f(x|\theta) = h(x)c(\theta)\exp\left(\sum_{i=1}^k \omega_i(\theta)t_i(x)\right)$$

then, from factorization theorem follows:

$T(\mathbf{X}) = \left(\sum_{j=1}^n t_1(X_j), \dots, \sum_{j=1}^n t_k(X_j)\right)$  is a sufficient statistics for  $\theta$ .

# Examples Sufficient Statistics

**Example:** The distribution of the sample under the Bernoulli distribution may be expressed as:

$$f(\mathbf{x}; \pi) = \pi^{\sum_i x_i} (1 - \pi)^{n - \sum_i x_i} = g(t(\mathbf{x}); \pi) h(\mathbf{x})$$

with  $t(\mathbf{x}) = \sum_i x_i$ ,  $g(t; \pi) = \pi^t (1 - \pi)^{n-t}$ , and  $h(\mathbf{x}) = 1$ . So, according to factorization theorem:  $Y = \sum_{i=1}^n X_i$  is a sufficient statistics for  $\pi$ .

# Examples Sufficient Statistics

**Example:** Let  $X_1, \dots, X_n$  be iid r.v. distributed as continuous uniform distribution on  $[0, \theta]$ . The probability distribution function of  $X_i$  for each  $i$  is:

$$f(x|\theta) = \begin{cases} \theta^{-1}, & 0 \leq x \leq \theta \\ 0, & \text{otherwise} \end{cases}$$

Thus the joint probability distribution function of  $X_1, \dots, X_n$  is

$$f(\underline{x}|\theta) = \begin{cases} \theta^{-n}, & 0 \leq x_i \leq \theta, \text{ for } i = 1, \dots, n \\ 0, & \text{otherwise} \end{cases}$$

# Examples Sufficient Statistics

**Example continued:** Since for each  $i$ ,  $x_i \leq \theta$ , we can write  $\max_i x_i \leq \theta$ , and let define  $T(x) = \max_i x_i$ :

$$f(\underline{x}|\theta) = \theta^{-n} \prod_i I_{[0,\theta]}(x_i)$$

$$f(\underline{x}|\theta) = \theta^{-n} I_{[\max(x_i),\infty]}(\theta)$$

and

$$g(t|\theta) = \begin{cases} \theta^{-n}, & t \leq \theta \\ 0, & \text{otherwise} \end{cases}$$

It can be easily verified that  $f(x|\theta) = g(T(x)|\theta)h(x)$  for all  $x$  and all  $\theta$ . So the statistics  $T(\mathbf{X}) = \max_i X_i$  is a sufficient statistics for  $\theta$ ,

# Examples Sufficient Statistics: Pareto

Let  $(X_1, \dots, X_n)$  be a random sample of i.i.d. random variables distributed as a Pareto distribution with parameters  $\alpha$  and  $x_m$

$$f(x|\alpha, x_m) = \frac{\alpha x_m^\alpha}{x^{\alpha+1}} \quad \text{for } x \geq x_m.$$

Find a sufficient statistics for  $\alpha$ ,

## Examples Sufficient Statistics: Pareto

$$\begin{aligned} f(x|\alpha, x_m) &= \prod_{i=1}^n \frac{\alpha x_m^\alpha}{x_i^{\alpha+1}} \\ &= \frac{\alpha^n x_m^{n\alpha}}{\prod_{i=1}^n x_i^{\alpha+1}} \\ &= g(t, \alpha)h(x) \end{aligned}$$

where  $t = \prod_{i=1}^n x_i$   $g(t, \alpha) = \alpha^n x_m^{n\alpha} t^{-(\alpha+1)}$  and  $h(x) = 1$ . By the factorization theorem,  $T(X) = \prod_{i=1}^n X_i$  is sufficient for  $\alpha$ .

# Examples Sufficient Statistics

Let  $(X_1, \dots, X_n)$  be a random sample of i.i.d. random variables distributed as a Beta distribution with parameters  $\alpha$  and  $\beta = 2$

$$f(x; \alpha) = \alpha(\alpha + 1)x^{\alpha+1}(1 - x) \quad 0 < x < 1 \quad \alpha > 0$$

- Prove that  $\sum \log x_i$  is a sufficient statistics for  $\alpha$ ,

# Examples Sufficient Statistics

$$f(x_1, \dots, x_n) = \alpha^n (\alpha + 1)^n \prod x_i^{\alpha+1} \prod (1 - x_i)$$

# Minimal Sufficient Statistics

- Many sufficient statistics may exist for the same parameter  $\theta$ .
- Obviously, also the sample  $X$  is a sufficient statistic for  $\theta$  as well as the order statistics  $Y_1 \leq Y_2 \leq Y_3 \leq \dots \leq Y_n$ . But they do not provide any data reduction.
- A sufficient statistic  $Y^* = T^*(\mathbf{X})$  is better than another sufficient statistic,  $Y = T(\mathbf{X})$  if  $Y^*$  achieves more data reduction while still retaining all the information on  $\theta$

# Minimal Sufficient Statistics

**Definition** The sufficient statistic that achieves the most data reduction is called **minimal sufficient statistic**.  
Formally, a sufficient statistic for  $\theta$ ,  $Y^* = t^*(\mathbf{X})$ , is a minimal sufficient statistic if it is a function of any other sufficient statistic  $Y = T(\mathbf{X})$

# Minimal Sufficient Statistics

In practice, the minimal sufficient statistics gives the greatest data reduction without loss of information about parameters. Here a characterization:

**Lehmann-Scheffe Theorem** A statistic  $T$  is minimal sufficient if the following property holds:

For any two sample points  $x$  and  $y$ ,  $f(x; \theta)/f(y; \theta)$  does not depend on  $\theta$  if and only if  $T(x) = T(y)$ .

**Corollary** Minimal sufficient statistic is not unique. But any two are in one-to-one correspondence, so are equivalent.

# Minimal sufficiency in exponential family

**Theorem:** For iid observations from an exponential family

$$f(x|\theta) = h(x)c(\theta)\exp\left(\sum_{i=1}^n \omega_i(\theta)t_i(x)\right)$$

the statistic

$$T(X) = \left(\sum_{j=1}^n T_1(X_j), \dots, \sum_{j=1}^n T_k(X_j)\right)$$

is minimal sufficient for  $\theta$

# Example

Let consider the pdf

$$f(x; \theta) = \frac{1}{\theta} \exp\left(1 - \frac{x}{\theta}\right) \quad 0 < \theta < x$$

Find a sufficient statistic for parameter  $\theta$ .

We know it is possible to find a sufficient statistic by applying the Factorization Theorem. Given the hypothesis of the random sampling we can factorize the given joint pdf as follows:

## Example

$$\begin{aligned} f(x_1, x_2, \dots, x_n; \theta) &= \prod_{i=1}^n \left( \frac{1}{\theta} \exp \left( 1 - \frac{x_i}{\theta} \right) l_{(\theta, \infty)}(x_i) \right) \\ &= \frac{1}{\theta^n} \exp \left( \sum_{i=1}^n \left( 1 - \frac{x_i}{\theta} \right) \right) \prod_{i=1}^n l_{(\theta, \infty)}(x_i) \\ &= \frac{1}{\theta^n} \exp \left( n - \frac{1}{\theta} \sum_{i=1}^n x_i \right) l_{(\theta, \infty)}(x_{(1)}) \end{aligned}$$

# Example

Therefore we can set:  $h(x) = 1$  and

$$g(T_1(x), T_2(x); \theta) = \frac{1}{\theta^n} \exp\left(n - \frac{1}{\theta} \sum_{i=1}^n x_i\right) I_{(\theta, \infty)}(x_{(1)})$$

where the two sufficient statistics are the sample sum and sample minimum  $T_1(x) = \sum_{i=1}^n x_i$  and  $T_2(x) = x_{(1)}$ .  $T_1(x)$  and  $T_2(x)$  are two jointly sufficient statistics for the parameter  $\theta$ . Thus, the dimension of a minimal sufficient statistic (two) is greater than the dimension of the parameter space (one).

## Example: Geometric distribution

Let consider the pdf

$$p(x; \pi) = \pi(1 - \pi)^{x-1}$$

We know it is possible to find a sufficient statistic by applying the Factorization Theorem:

$$\begin{aligned} p(x_1, x_2, \dots, x_n; \pi) &= \prod_{i=1}^n \frac{\pi}{1 - \pi} (1 - \pi)^{x_i} \\ &= \left( \frac{\pi}{1 - \pi} \right)^n \exp \left( \sum_i x_i \log(1 - \pi) \right) \end{aligned}$$

So that sample sum,  $\sum_i x_i$ , is sufficient and complete statistics.

## Example: Uniform distribution

Suppose that  $X_1, X_2, \dots, X_n$  form a random sample from a uniform distribution on the interval  $[\theta - 1, \theta + 1]$ , with  $-\infty < \theta < \infty$ .

$$f(x|\theta) = 1 \quad \theta - 1 \leq x \leq \theta + 1$$

$$f(x_i|\theta) = 1 \quad \theta - 1 \leq x_i \leq \theta + 1$$

$$f(x_1, \dots, x_n|\theta) = 1 \quad \theta \leq \min(x_i) + 1 \text{ and } \theta \geq \max(x_i) - 1$$

$f(x_1, \dots, x_n|\theta) = f(y_1, \dots, y_n|\theta)$  is independent of  $\theta$  if and only if  $\min(x_i) = \min(y_i)$  and  $\max(x_i) = \max(y_i)$ , therefore  $T(X) = (\min(X_i), \max(X_i))$  is minimal sufficient.

## Exercises:

Assume that the random variables  $X_1, X_2, \dots, X_n$  form a random sample of size  $n$  from the distribution specified, and show that the statistic  $T$  specified is a sufficient statistic for the parameter:

**Exercise 1:** A normal distribution for which the mean  $\mu$  is known and the variance  $\sigma^2$  is unknown;  $T = \sum_{i=1}^n (X_i - \mu)^2$ .

**Exercise 2:** A gamma distribution with parameters  $\alpha$  and  $\beta$ , where the value of  $\beta$  is known and the value of  $\alpha$  is unknown ( $\alpha$ );  $T = \prod_{i=1}^n X_i$

**Exercise 3:** A uniform distribution on the interval  $[a, b]$ , where the value of  $a$  is known and the value of  $b$  is unknown ( $b > a$ );  $T = \max(X_1, \dots, X_n)$ .

**Exercise 4:** A uniform distribution on the interval  $[a, b]$ , where the value of  $b$  is known and the value of  $a$  is unknown ( $b > a$ );  $T = \min(X_1, \dots, X_n)$ .

## Exercises:

**Exercise 5:** Suppose that  $X_1, X_2, \dots, X_n$  form a random sample from a gamma distribution with parameters  $\alpha > 0$  and  $\beta > 0$ , and the value of  $\beta$  is known. Show that the statistic  $T = \sum_{i=1}^n \log X_i$  is a sufficient statistic for the parameter  $\alpha$ .

**Exercise 6:** The Pareto distribution has density function:

$$f(x|x_0; \theta) = \theta x_0^\theta x^{-\theta-1}, \quad x \geq x_0, \quad \theta > 1$$

Assume that  $x_0 > 0$  is given and that  $X_1, X_2, \dots, X_n$  is an i.i.d. sample. Find a sufficient statistic for *theta* by

- (a) using factorization theorem,
- (b) using the property of exponential family. Are they the same? If not, why are both of them sufficient?

## Example of in-sufficiency

Suppose that  $X_1$  and  $X_2$  are two independent Bernoulli random variables with parameter  $p$ ,  $0 < p < 1$ . Show that the statistics  $T = X_1 - X_2$  is not a sufficient statistics for  $p$ .

## Example of in-sufficiency

- $(X_1, X_2, \dots, X_n)$  iid  $Poisson(\lambda)$   $T = X_1 - X_2$  is not sufficient
- $(X_1, X_2, \dots, X_n)$  iid pmf  $f(x; \theta)$   $T = (X_1, X_2, \dots, X_{n-1})$  is not sufficient.